

Please adhere to the homework rules as given in the Syllabus.

**1. Trapezoidal Distribution.** Consider the following probability density function.

$$f(x) = \begin{cases} \frac{x+1}{5}, & -1 \leq x < 0 \\ \frac{1}{5}, & 0 \leq x < 4 \\ \frac{5-x}{5}, & 4 \leq x \leq 5 \\ 0, & \text{otherwise} \end{cases}$$

- a) Sketch the PDF.
  
  
  
  
  
  
  
  
- b) Show that the area under the curve is equal to 1 using geometry.
  
  
  
  
  
  
  
  
- c) Show that the area under the curve is equal to 1 by integrating. (You will have to split the integral into pieces).
  
  
  
  
  
  
  
  
- d) Find  $P(X < 3)$  using geometry.
  
  
  
  
  
  
  
  
- e) Find  $P(X < 3)$  with integration.

**2.** Let  $X$  have the following PDF, for  $\theta > 0$ .

$$f(x) = \begin{cases} \frac{c}{\theta^2}(\theta - x), & 0 < x < \theta \\ 0, & \text{otherwise} \end{cases}$$

a) Find the value of  $c$  which makes  $f(x)$  a valid PDF. Sketch the PDF.

b) Find the mean and standard deviation of  $X$ .

c) Find  $P\left(\frac{\theta}{10} < X < \frac{3\theta}{5}\right)$ .

**3. PDF to CDF.** For each of the following PDFs, find and sketch the CDF. The Median ( $M$ ) of a continuous distribution is defined by the property  $F(M) = 0.5$ . Use the CDF to find the Median of each distribution.

a) (Special case of Beta Distribution)

$$f(x) = 1.5\sqrt{x}, \quad 0 < x < 1$$

b) (Special case of Pareto Distribution)

$$f(x) = \frac{1}{x^2}, \quad x > 1$$

**4. CDF to PDF.** For each of the following CDFs, find the PDF.

a)

$$F(x) = \begin{cases} 0, & x < 0 \\ 1 - (1 - x^2)^2, & 0 \leq x < 1 \\ 1, & x \geq 1 \end{cases}$$

b)

$$F(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-x^2}, & x \geq 0 \end{cases}$$

## 5. Expected Values.

- a) Continuous uniform distribution: Let  $X$  have PDF  $f(x) = 1/\pi$  for  $0 < x < \pi$ . Find  $E(\pi^2 \sin(X))$ .
- b) Let  $X$  have PDF,  $f(x) = 2x$  for  $0 < x < 1$ . Use integration by parts or tabular integration to find  $E(Xe^X)$ .
- c) Exponential distribution: Let  $X$  have PDF  $f(x) = \lambda e^{-\lambda x}$  for  $x > 0$  where  $\lambda > 0$  is a constant. Find  $E(a^X)$  where  $a > 0$ . Does  $E(a^X)$  exist for all values of  $a > 0$ ? If not, give a condition on  $a$  for which the expected value exists. *Hint:*  $a^x e^{-\lambda x} = (ae^{-\lambda})^x$

**6. Waiting Times.** Suppose that the time it takes for Jules to get his food at Big Kahuna Burger is a random variable  $X$  with the following CDF (*Note: This is an Exponential distribution with an expected waiting time of 1 minute.*).

$$F_X(x) = 1 - e^{-x}, \quad x > 0$$

- a) Find the probability that Jules waits at least 1 minute before getting his tasty burger.

b) Given that Jules has already waited 5 minutes, what is the probability he must wait at least 1 more minute before getting his tasty burger. Compare this answer to part a). *Hint: Use the definition of Conditional Probability to find  $P(X > 6|X > 5)$ .*

c) Jules is happy as long as he gets his burger within 2 minutes. Let  $Y = 1$  if Jules is happy and  $Y = 0$  otherwise. What is the distribution of  $Y$ ? Also give the mean and variance of  $Y$ .

c) Jules visits Big Kahuna Burger every monday for a year. Each time he records his waiting time  $X_i$  and the corresponding random variable  $Y_i$  defined above for  $i = 1, 2, \dots, 52$ . Assume each visit is independent, and let  $Z = \sum_{i=1}^{52} Y_i$ . What is the distribution, mean and variance of  $Z$ ?

**7. Challenge Problem:** The Moment Generating Function (MGF) of a random variable  $X$  is defined as

$$M_X(t) = E(e^{Xt})$$

Inside the expected value,  $t$  can be regarded as a constant. However,  $M_X()$  is a function of this  $t$ . (*Note: If you are familiar with Laplace Transforms, note that for continuous  $X$ , the MGF is essentially the Laplace Transform of the PDF*).

- a) Find the Moment Generating Function of  $X \sim Exp(\lambda)$ . You will need to look the PDF  $f(x)$  up in Table 1. *Hint: With some manipulation, you can use the result from problem 5c.*

Moment Generating Functions are useful for a variety of reasons (the most important of which will appear as a future challenge problem). As the name suggests, they can be useful for finding moments of the distribution. Specifically, the following identity holds for any non-negative integer  $k$ .

$$E(X^k) = M^{(k)}(0)$$

*Notation:  $M_X^{(k)}(0)$  is the  $k^{th}$  derivative of  $M_X(t)$  with respect to  $t$ , evaluated at 0.*

- b) Use this identity to find  $E(X)$  and  $E(X^2)$  for  $X \sim Exp(\lambda)$ . You should get the same answer as with the usual method.

I should note that the MGF does not exist for some distributions, in which case this method cannot be used. Now let's try to see why the MGF works!

- c) Expand  $M_X(t) = E(e^{Xt})$  by writing  $e^{Xt}$  as a Taylor Series expansion. Now use linearity of expectation. Explain (no need for formal proof) why the identity holds.