

# Nonparametric Density Estimation under Total Positivity

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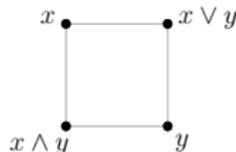
May 24, 2021

# Totally Positive Distributions

- A distribution with density  $p$  on  $\mathcal{X} \subseteq \mathbb{R}^d$  is *multivariate totally positive of order 2* (or  $MTP_2$ ) if

$$p(x)p(y) \leq p(x \wedge y)p(x \vee y) \quad \text{for all } x, y \in \mathcal{X},$$

where  $x \wedge y$  and  $x \vee y$  are the componentwise minimum and maximum.

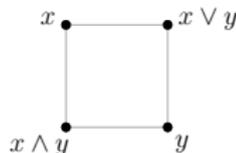


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where  $x \wedge y$  and  $x \vee y$  are the componentwise minimum and maximum.



- $MTP_2$  is the same as *log-supermodular*:

$$\log(p(x)) + \log(p(y)) \leq \log(p(x \wedge y)) + \log(p(x \vee y)) \quad \text{for all } x, y \in \mathcal{X}.$$

# Properties of $MTP_2$ distributions

**Theorem** (Fortuin Kasteleyn Ginibre, 1971)

*$MTP_2$  implies positive association.*

## Definition

A random vector  $X$  taking values in  $\mathbb{R}^d$  is *positively associated* if for any non-decreasing functions  $\phi, \psi : \mathbb{R}^d \rightarrow \mathbb{R}$

$$\text{cov}(\phi(X), \psi(X)) \geq 0.$$

**Theorem** (Karlin and Rinot 1980; Lebowitz 1972)

If  $X = (X_1, \dots, X_d)$  is  $MTP_2$ , then

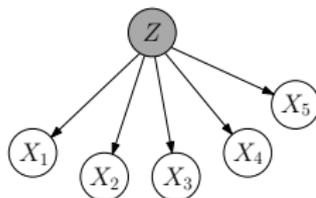
- (i) any marginal distribution is  $MTP_2$ ,
- (ii) any conditional distribution is  $MTP_2$ ,
- (iii)  $X$  has the marginal independence structure

$$X_i \perp\!\!\!\perp X_j \iff \text{cov}(X_i, X_j) = 0.$$



# Examples of $MTP_2$ distributions

- The joint distribution of observed variables influenced by one hidden variable (in the Gaussian and binary case)



- Many models imply  $MTP_2$ :
  - Ferromagnetic Ising models
  - Order statistics of i.i.d. variables
  - Brownian motion tree models
  - Gaussian/binary (latent) tree models (e.g. single factor analysis models)

## Examples of $MTP_2$ distributions

A Gaussian random variable  $X \sim \mathcal{N}(\mu, \Sigma)$  is  $MTP_2$  whenever  $\Sigma^{-1}$  is an M-matrix, i.e. its off-diagonal entries are nonpositive

$$(\Sigma^{-1})_{ij} \leq 0, \quad \forall i \neq j.$$

Very rare: out of 100,000 uniformly sampled  $5 \times 5$  correlation matrices none were  $MTP_2$ . [Fallat et al., 2016].

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Data: grades of 88 students in 5 math subjects

[Mardia, Kent and Bibby, 1979, Fallat et al., 2016]

$$S = \begin{pmatrix} \text{mechanics} & \text{vectors} & \text{algebra} & \text{analysis} & \text{statistics} \\ 305.7680 & 127.2226 & 101.5794 & 106.2727 & 117.4049 \\ 127.2226 & 172.8422 & 85.1573 & 94.6729 & 99.0120 \\ 101.5794 & 85.1573 & 112.8860 & 112.1134 & 121.8706 \\ 106.2727 & 94.6729 & 112.1134 & 220.3804 & 155.5355 \\ 117.4049 & 99.0120 & 121.8706 & 155.5355 & 297.7554 \end{pmatrix} \begin{matrix} \text{mechanics} \\ \text{vectors} \\ \text{algebra} \\ \text{analysis} \\ \text{statistics} \end{matrix}$$

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$$S^{-1} = 10^{-3} \cdot \begin{pmatrix} \text{mechanics} & \text{vectors} & \text{algebra} & \text{analysis} & \text{statistics} \\ 5.2446 & -2.4351 & -2.7395 & \mathbf{0.0116} & -0.1430 \\ -2.4351 & 10.4268 & -4.7078 & -0.7928 & -0.1660 \\ -2.7395 & -4.7078 & 26.9548 & -7.0486 & -4.7050 \\ \mathbf{0.0116} & -0.7928 & -7.0486 & 9.8829 & -2.0184 \\ -0.1430 & -0.1660 & -4.7050 & -2.0184 & 6.4501 \end{pmatrix} \begin{matrix} \text{mechanics} \\ \text{vectors} \\ \text{algebra} \\ \text{analysis} \\ \text{statistics} \end{matrix}$$

# Examples of MTP<sub>2</sub> distributions

Data: 2016 monthly correlations of global stock markets

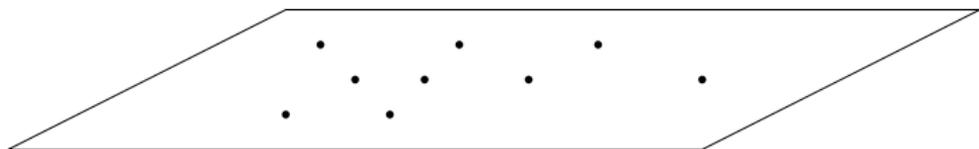
[*InvestmentFrontier.com*; Agrawal, Roy, and Uhler, 2019]

$$S = \begin{pmatrix} \text{Nasdaq} & \text{Canada} & \text{Europe} & \text{UK} & \text{Australia} \\ 1.000 & 0.606 & 0.731 & 0.618 & 0.613 \\ 0.606 & 1.000 & 0.550 & 0.661 & 0.598 \\ 0.731 & 0.550 & 1.000 & 0.644 & 0.569 \\ 0.618 & 0.661 & 0.644 & 1.000 & 0.615 \\ 0.613 & 0.598 & 0.569 & 0.615 & 1.000 \end{pmatrix} \begin{matrix} \text{Nasdaq} \\ \text{Canada} \\ \text{Europe} \\ \text{UK} \\ \text{Australia} \end{matrix}$$

$$S^{-1} = \begin{pmatrix} \text{Nasdaq} & \text{Canada} & \text{Europe} & \text{UK} & \text{Australia} \\ 2.629 & -0.480 & -1.249 & -0.202 & -0.490 \\ -0.480 & 2.109 & -0.039 & -0.790 & -0.459 \\ -1.249 & -0.039 & 2.491 & -0.675 & -0.213 \\ -0.202 & -0.790 & -0.675 & 2.378 & -0.482 \\ -0.490 & -0.459 & -0.213 & -0.482 & 1.992 \end{pmatrix} \begin{matrix} \text{Nasdaq} \\ \text{Canada} \\ \text{Europe} \\ \text{UK} \\ \text{Australia} \end{matrix}$$

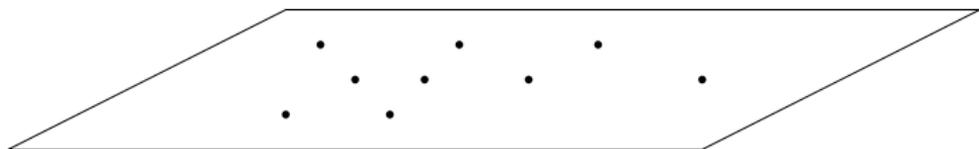
# Density estimation

Given i.i.d. samples  $X = \{x_1, \dots, x_n\} \subset \mathbb{R}^d$  from an unknown distribution on  $\mathbb{R}^d$  with density  $p$ , can we estimate  $p$ ?



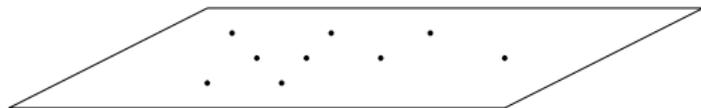
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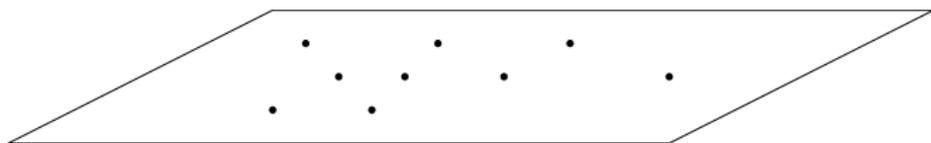
- non-parametric: assume that  $p$  lies in a non-parametric family, e.g. impose shape-constraints on  $p$  (convex, log-concave, monotone, etc.)
  - infinite-dimensional problem
  - need constraints that are:
    - strong enough so that there is no spiky behavior
    - weak enough so that function class is large

# Shape-constrained density estimation



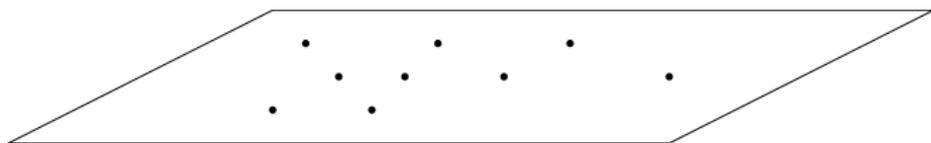
- monotonically decreasing densities: [Grenander 1956, Rao 1969]
- convex densities: [Anevski 1994, Groeneboom, Jongbloed, and Wellner 2001]
- **log-concave** densities: [Cule, Samworth, and Stewart 2008, Dümbgen et al. 2009, . . .]
- generalized additive models with shape constraints: [Chen and Samworth 2016]
  
- **totally positive (and log-concave) densities**

## Maximum Likelihood Estimation under $MTP_2$



Given i.i.d. samples  $X = \{x_1, \dots, x_n\} \subset \mathbb{R}^d$  with weights  $w = (w_1, \dots, w_n)$  (where  $w_1, \dots, w_n \geq 0$ ,  $\sum w_i = 1$ ) from a distribution  $p$  on  $\mathbb{R}^d$ , can we estimate  $p$ ?

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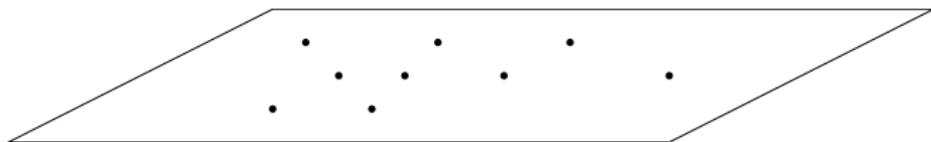


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We would like to

$$\begin{aligned} & \text{maximize}_p \quad \sum_{i=1}^n w_i \log(p(x_i)) \\ & \text{s.t.} \quad p \text{ is an } MTP_2 \text{ density.} \end{aligned}$$

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**Problem:** likelihood is unbounded,  $MTP_2$  constraint is too weak for MLE to exist.



## Maximum Likelihood Estimation under $MTP_2$

To ensure that the likelihood function is bounded, we impose the condition that  $p$  is log-concave.

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- Log-concavity is a natural assumption because it ensures the density is continuous and includes many known families of parametric distributions.
- Log-concave families:
  - Gaussian;
  - Uniform( $a, b$ );
  - Gamma( $k, \theta$ ) for  $k \geq 1$ ;
  - Beta( $a, b$ ) for  $a, b \geq 1$ .
- Maximum likelihood estimation under log-concavity is a well-studied problem (Cule et al. 2008, Dümbgen et al. 2009, Schuhmacher et al. 2010, ...).

# Maximum Likelihood Estimation under Log-Concavity

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## Theorem (Cule, Samworth and Stewart 2008)

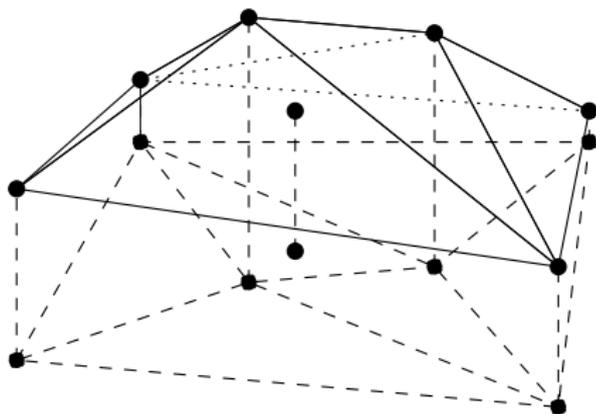
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# Maximum Likelihood Estimation under Log-Concavity

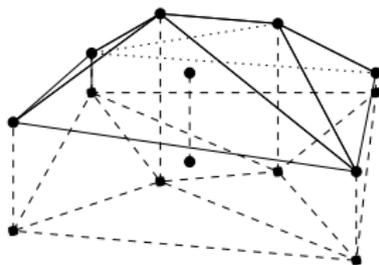
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- With probability 1, a log-concave maximum likelihood estimator  $\hat{p}$  exists and is unique as long as there are at least  $d + 1$  samples.
- Moreover,  $\log(\hat{p})$  is a 'tent-function' supported on the convex hull of the data  $P(X) = \text{conv}(x_1, \dots, x_n)$ .



## Optimizing over Tent Functions

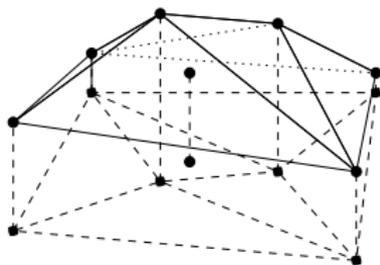


Given points  $X = \{x_1, \dots, x_n\}$  and heights  $y = (y_1, \dots, y_n) \in \mathbb{R}^n$ , the *tent function*

$$h_{X,y} : \mathbb{R}^d \rightarrow \mathbb{R}$$

is the smallest concave function such that  $h_{X,y}(x_i) \geq y_i$  for all  $i$ . Thus,  $\hat{p} = \exp(h_{X,y})$  for some  $y$ .

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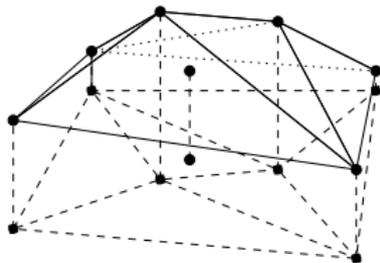
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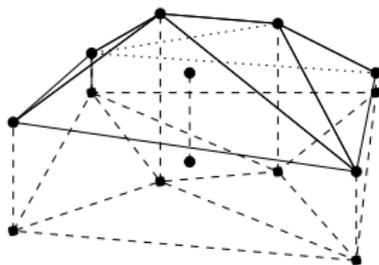
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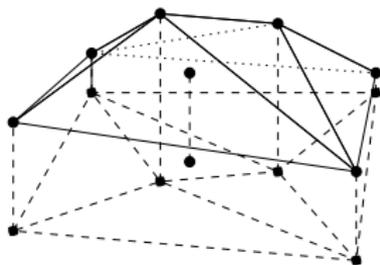
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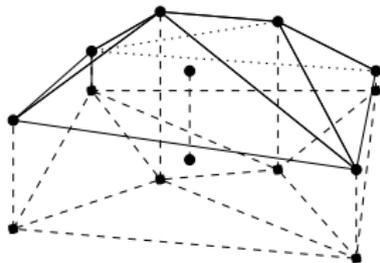
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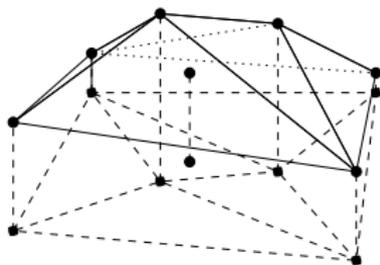
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$$\text{max}_{y \in \mathbb{R}^n} \sum_{i=1}^n w_i y_i - \int \exp(h_{X,y}(t)) dt$$

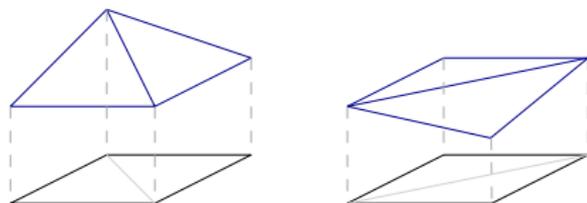
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# Maximum Likelihood Estimation under Log-concavity and $MTP_2$

Questions:

1. Does the MLE under log-concavity and  $MTP_2$  exist with probability 1 and, if so, is it unique?
2. What is the shape of the MLE under log-concavity and  $MTP_2$ ?
  - 2.1 What is the support of the MLE?
  - 2.2 Is the MLE always exp(tent function)?
3. Which tent functions are allowed?
4. Can we compute the MLE?
5. Sample complexity?



# Existence and Uniqueness of the MLE

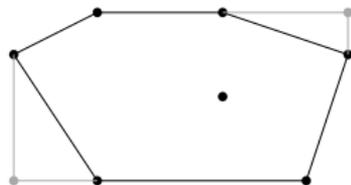
## Theorem (R., Sturmfels, Tran, Uhler, 2018)

- *The maximum likelihood estimator under log-concavity and  $MTP_2$  exists and is unique with probability 1 as long as there are at least 3 samples.*
- *Furthermore, the MLE is a consistent estimator.*

Proof uses convergence properties for log-concave distributions, and does not shed light on the shape of the MLE.

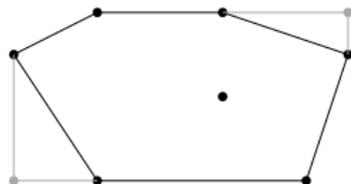
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Support of MLE = "Min-max convex hull" of  $X$ .



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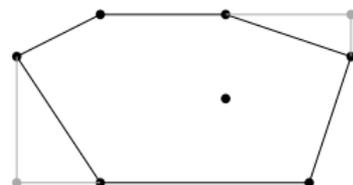
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- $MMconv(X)$  = smallest *min-max* closed & convex set containing  $X$ .

# The Support of the MLE

Support of MLE = "Min-max convex hull" of  $X$ .

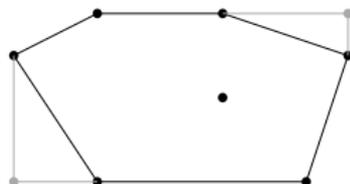


- $MMconv(X)$  = smallest *min-max closed & convex* set containing  $X$ .
- $MM(X)$  = smallest *min-max closed* set  $S$  containing  $X$ , i.e.  
 $x, y \in S \Rightarrow x \wedge y, x \vee y \in S$

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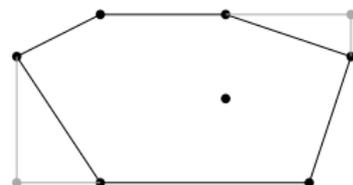


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## Lemma

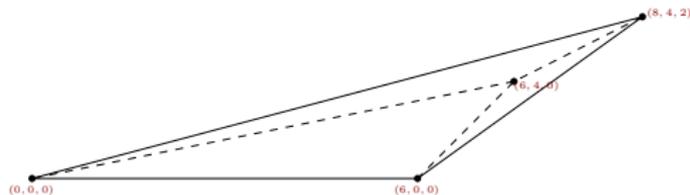
Let  $X = \{x_1, \dots, x_n\}$ . If  $X \subseteq \mathbb{R}^2$  or  $X \subseteq \prod_{i=1}^d \{a_i, b_i\}$ , then,

$$MMconv(X) = conv(MM(X)).$$

$\implies$  we can compute  $MMconv(X)$  for  $X \subseteq \mathbb{R}^2$  and  $X \subseteq \prod_{i=1}^d \{a_i, b_i\}$ .

# The Min-Max Convex Hull

Now, consider  $X = \{(0, 0, 0), (6, 0, 0), (6, 4, 0), (8, 4, 2)\} \subseteq \mathbb{R}^3$ .



It turns out that

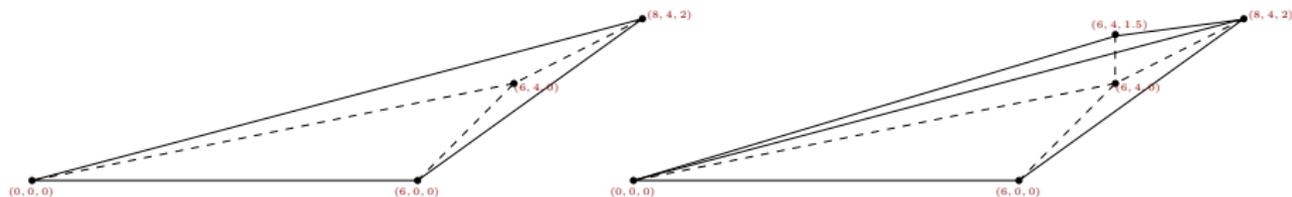
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But

$\text{conv}(\text{MM}(X))$  is **not** min-max closed!

This is because:

$$\left(6, 4, \frac{3}{2}\right) = \max\left\{(6, 4, 0), \left(6, 3, \frac{3}{2}\right)\right\} \notin \text{conv}(\text{MM}(X)).$$

Therefore,

$$\text{conv}(\text{MM}(X)) \subsetneq \text{MMconv}(X).$$

# Computing $MMconv(X)$

## Theorem (The 2-D Projections Theorem)

For any finite subset  $X \subseteq \mathbb{R}^d$ ,

$$MMconv(X) = \bigcap_{1 \leq i < j \leq d} \pi_{ij}^{-1}(\text{conv}(MM(\pi_{ij}(X)))).$$

$$\begin{aligned} \pi_{ij} : \mathbb{R}^d &\rightarrow \mathbb{R}^2, \\ x &\mapsto (x_i, x_j). \end{aligned}$$

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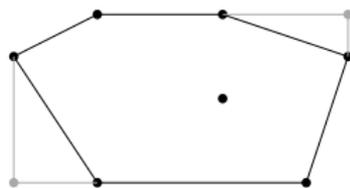
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## Corollary (Queyranne and Tardella, 2006)

A convex polytope  $C$  in  $\mathbb{R}^d$  is min-max closed if and only if it is defined by a finite collection of bimonotone linear inequalities.

A linear inequality is *bimonotone* if it has the form

$$a_i x_i + a_j x_j + b \leq 0, \quad \text{where } a_i a_j \leq 0.$$



# Back to Log-concave and $MTP_2$ Maximum Likelihood Estimation

1. Does the MLE under log-concavity and  $MTP_2$  exist with probability 1 and, if so, is it unique? **Yes.**
2. What is the shape of the MLE under log-concavity and  $MTP_2$ ?
  - 2.1 What is the support of the MLE?  **$MMconv(X)$ ; We can compute it.**
  - 2.2 Is the MLE always exp(tent function)?
3. Which tent functions are allowed?
4. Can we compute the MLE?
5. Sample complexity?

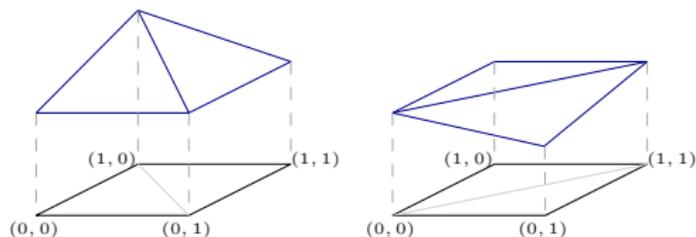
# Supermodular Tent Functions

Recall that  $p = \exp(h)$  is **MTP**<sub>2</sub> if and only if  $h$  is **supermodular**, i.e.

$$h(x) + h(y) \leq h(x \wedge y) + h(x \vee y), \quad \text{for all } x, y \in \mathbb{R}^d.$$

**Theorem** (R., Sturmfels, Tran, Uhler, 2018)

Let  $X \subset \mathbb{R}^d$  be a finite set of points. A tent function  $h$  is supermodular if and only if all of the walls of the subdivision  $h$  induces are **bimonotone**.



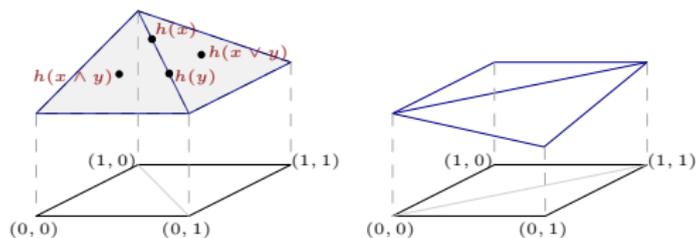
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## Remark

To find the best  $h_{X,y}$ , need to optimize over heights  $y$  that induce bimonotone subdivisions.

- In general not convex.
- Example:  $X = \{0, 1\} \times \{0, 1\} \times \{0, 1, 2\}$ .

# Maximum Likelihood Estimation under Log-concavity and $MTP_2$

1. Does the MLE under log-concavity and  $MTP_2$  exist with probability 1 and, if so, is it unique? **Yes.**
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# When is the MLE the exponential of a tent function?

Theorem (R., Sturmfels, Tran, Uhler 2018)

If  $X \subseteq \mathbb{R}^2$  or  $X \subseteq \prod_{i=1}^d \{a_i, b_i\}$ , then the MLE is the exponential of a tent function

$$p^* = \exp(h_{\tilde{X}, y}),$$

where  $\tilde{X} = MM(X)$ .

# Finding the MLE

Theorem (R., Sturmfels, Tran, Uhler 2018)

If  $X \subseteq \mathbb{R}^2$  or  $X \subseteq \prod_i \{a_i, b_i\}$ , then the set of heights for which  $\exp(h_{\tilde{X}, y})$  is  $MTP_2$  is a convex polytope  $\mathcal{S}$ .

Therefore, we can use, e.g. the conditional gradient method, to find  $y^*$ .

$$\begin{aligned} & \underset{y}{\text{maximize}} && \sum_{i=1}^n w_i y_i - \int \exp(h_{\tilde{X}, y}) \\ & \text{s.t.} && y \in \mathcal{S}. \end{aligned}$$

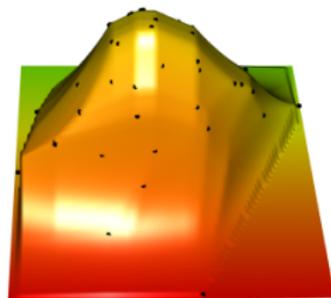
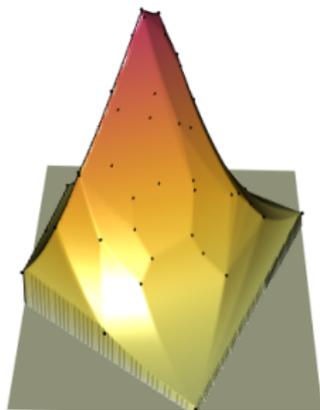
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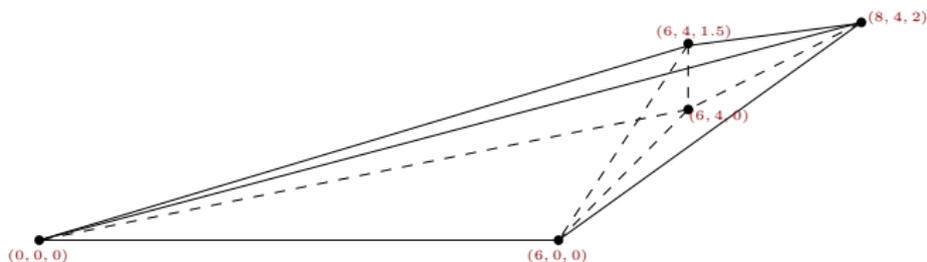
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# What is the shape of the MLE in the general case?

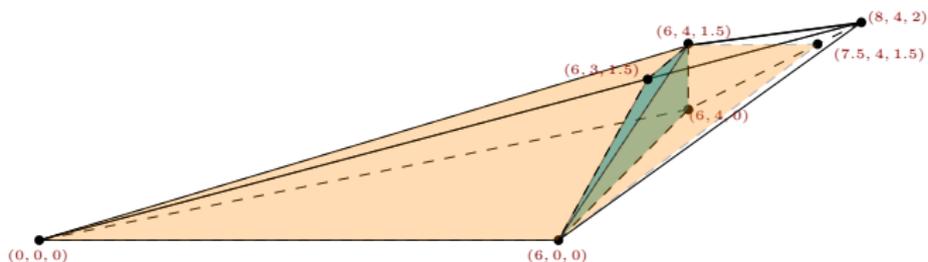
Let  $X = \{(0, 0, 0), (6, 0, 0), (6, 4, 0), (8, 4, 2), (6, 4, \frac{3}{2})\}$ ,  $w = \frac{1}{28}(15, 1, 1, 1, 10)$ . The log-concave MLE is not  $MTP_2$ .



The  $MTP_2$  and log-concave MLE is  $\exp(\text{tent function})$  on  $X \cup \{(6, 3, \frac{3}{2}), (7.5, 4, \frac{3}{2})\}$  with subdivision as above.

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# Sample complexity for log-concave density estimation

- $\mathcal{F}_d$  = set of log-concave and MTP<sub>2</sub> densities in  $\mathbb{R}^d$
- $h^2(f, g) = \int (\sqrt{f} - \sqrt{g})^2$  Hellinger distance
- $\hat{p}_n$  = log-concave MLE from  $n$  samples

## Theorem (Kim and Samworth, 2014)

(a). *Upper bound for  $d = 2, 3$ :*

$$\sup_{p_0 \in \mathcal{F}_2} \mathbb{E}_{p_0} [h^2(p_0, \hat{p}_n)] = \mathcal{O}(n^{-\frac{2}{d+1}} \log n).$$

(b). *Lower bound for  $d \geq 2$ :*

$$\inf_{\tilde{p}_n} \sup_{p_0 \in \mathcal{F}_2} \mathbb{E}_{p_0} [h^2(p_0, \tilde{p}_n)] \geq c_2 n^{-\frac{2}{d+1}}.$$

*MLE is minimax optimal up to log factors for  $d = 2, 3$ .*

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## Theorem (Carpenter et al., 2018)

When  $d \geq 4$ :

(a). Upper bound:

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## Problem

When  $d \geq 3$ :

(a). Upper bound?

$$\sup_{p_0 \in \mathcal{F}_d} \mathbb{E}_{p_0}[h^2(p_0, \hat{p}_n)] = \mathcal{O}(n^{-\frac{2}{3}} \log n)$$

(b). Lower bound:

$$\inf_{\tilde{p}_n} \sup_{p_0 \in \mathcal{F}_d} \mathbb{E}_{p_0}[h^2(p_0, \tilde{p}_n)] \geq c_2 \text{poly}(d) n^{-\frac{2}{3}}$$

No curse of dimensionality?

# Conclusion

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- Rates for  $d \geq 3$ ?
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- Min-max closure of sample,  $MM(X)$ , convolved with standard Gaussian is  $MTP_2$ .  
[with Ali Zartash, 2019]
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[with Ali Zartash, 2019]
- Consistency and rates?
  
- Test statistics for  $MTP_2$  distributions?
- Multivariate total positivity of higher orders (i.e.  $MTP_k$ )?

Thank you!

# References



M. Cule, R. Samworth, M. Stuart. *Maximum likelihood estimation of a multidimensional log-concave density*. *Statistical Methodology Series B* 72:5 (2010), 545-607



J.-C. Hütter, C. Mao, P. Rigollet, and E. Robeva. *Optimal Rates for Estimation of Two-dimensional Totally Positive Distributions*.



**E. Robeva**, B. Sturmfels, and C. Uhler. *Geometry of Log-concave Density Estimation*. arXiv:1704.01910



**E. Robeva**, B. Sturmfels, N. Tran, and C. Uhler. *Maximum Likelihood Estimation for Totally Positive Log-Concave Densities*. arXiv:1806.10120



**E. Robeva** and M. Sun. *Bimonotone Subdivisions of Point Configurations in the Plane*. In preparation



**E. Robeva**, C. Uhler, and Y. Wang. *Rates of Estimation of Log-Concave and Totally Positive Densities*. In preparation



A. Zartash and **E. Robeva**. *Kernel density estimation for totally positive random vectors*. arXiv:1910.02345